

types aggregate in QTL regions, combined with physiological profiling, provides a novel approach to facilitate the positional cloning of genes underlying cardiovascular function and hypertension. While strategies for accomplishing these goals will continue to evolve, this first attempt at developing knowledge at a systems biology level sets the stage for future improvements and provides investigators with a powerful set of tools for discovery.

References and Notes

1. E. R. Bleecker, D. S. Postma, D. A. Meyers, *Am. J. Respir. Crit. Care Med.* **156**, S113 (1997).
2. C. Julier *et al.*, *Hum. Mol. Genet.* **6**, 2077 (1997).
3. J. Krushkal *et al.*, *Circulation* **99**, 1407 (1999).
4. C. L. Hanis *et al.*, *Nature Genet.* **13**, 161 (1996).
5. J. A. Todd, *Proc. Natl. Acad. Sci. U.S.A.* **92**, 8560 (1995).
6. A user interface to the complete data set is found at <http://brc.mcw.edu/phyprf>.
7. The total genome scan was carried out with an average 10 cM spacing of markers following phenotyping with 239 measured or derived traits. After testing for normalcy, 166 traits were analyzed in a parametric genome scan using MAPMAKER/QTL (12, 16). The remaining 73 traits were analyzed using the nonparametric mapping algorithm (MAPMAKER/QTL version 1.9b). (Phenotyping protocol, conscious): All blood pressure measurements were made with the animals unrestrained in their home cage as described previously (17). Data were collected at a rate of 100 Hz and reduced to 1-min averages, except for time series analysis where they were 1-s averages. The day-night light cycle for all rats ran from 2:00 a.m. (lights on) to 2:00 p.m. (lights off) throughout the study. BP1 (High Salt Day 1): Baseline measurements of systolic, diastolic, and mean arterial pressure and heart rate were measured from 9:00 a.m. to noon. BP2 (High Salt Day 2 – inactive versus active): Repeat morning recording ("inactive phase," lights on), and record 4 hours during the dark cycle (2:00 p.m. to 6:00 p.m. "active phase"). Urine was collected for 24 hours for measurement of volume, sodium, potassium, protein, and creatinine. All BP data this day was collected for time-series analysis. BP3 (High Salt Day 3): Repeat morning recording. Following the recording period, a blood sample (500 μ l) was drawn for determination of creatinine, plasma renin activity, plasma protein, and hematocrit. Furosemide challenge: Following blood draw, an ip injection of furosemide (10 mg/kg) was given to salt deplete the animals, and the diet switched to a 0.4% low salt. BP4 (Stress test): Following a control period, an alerting stimulus (2 mA for 0.3 s) was delivered twice with a 5-min interval during pressure recording. Change in mean arterial pressure, the time to peak, and the time to 90% recovery were determined. BP5 (Low Salt): Repeat 3-hour morning recording of blood pressure in the salt depleted state. Following the recording period, a 1.0-ml blood sample was taken for the measurement of plasma renin activity, triglycerides, total cholesterol, HDL, creatinine, hematocrit, and white blood cell count. (Phenotyping protocol, anesthetized): Rats were anesthetized with ketamine (30 mg/kg; intramuscular) and Inactin (50 mg/kg; intraperitoneal). Catheters were implanted in the femoral artery and vein and an electromagnetic flow probe was placed on the left renal artery via a midline incision. An iv infusion (50 μ l/min) of isotonic saline containing 1% bovine serum albumin replaced fluid losses. After a 45-min equilibration, arterial blood pressure and renal blood flow were measured during a 15-min control period. Renal and peripheral vascular responses to 5-min iv infusions of angiotensin II (20, 100, 200 ng/kg/min) and norepinephrine (0.5, 1, 3 μ g/kg/min) were determined. Following recovery of pressure to baseline values, renal vascular and systemic arterial responses to two successive doses of acetylcholine (ACh) (0.1 and 0.2 μ g/kg/min) were measured. L-NAME (5 mg/kg) was then administered

as an iv bolus to determine the contribution of nitric oxide to basal renal vascular tone. After 10 min of equilibration, a repeat infusion of the same two doses of ACh were administered to test for the degree of blockade of the synthesis of nitric oxide produced by L-NAME. (Morphometric measurements): Heart and kidneys were removed, stripped of surrounding tissue, and weighed to assess the degree of cardiac and renal hypertrophy. (Histology): The right kidney was immersion-fixed in 10% buffered formalin and embedded in paraffin, and prepared sections stained with H&E and PAS were evaluated for mean glomerular diameter and the degree of focal glomerulosclerosis.

8. B. R.Thumma *et al.*, *J. Exp. Bot.* **52**, 203 (2001).
9. L. L. Miner *et al.*, *Psychopharmacology* **117**, 62 (1995).
10. K. G. Becker *et al.*, *Proc. Natl. Acad. Sci. U.S.A.* **95**, 9979 (1998).
11. J. Nadeau, W. Frankel, *Nature Genet.* **25**, 381 (2000).
12. J. P. Rapp, *Physiol. Rev.* **80**, 135 (2000).
13. A. C. Guyton *et al.*, *Annu. Rev. Physiol.* **34**, 13 (1972).

14. E. S. Lander *et al.*, *Genomics* **1**, 174 (1987).
15. J. Loscalzo, G. Welch, *Progr. Cardiovasc. Dis.* **38**, 87 (1995).
16. N. K. Hollenberg *et al.*, *Medicine* **57**, 167 (1978).
17. A. W. Cowley Jr. *et al.*, *Physiol. Genomics* **2**, 107 (2000).
18. M. Stoll, A. W. Cowley Jr., A. S. Greene, data not shown.
19. E. S. Lander, L. Kruglyak, *Nature Genet.* **11**, 241 (1995).
20. We thank the Bioinformatics Research Center at Medical College of Wisconsin for the development of the database, tools, and Web site; M. Granados, M. Nobrega, M. Shiozawa, and M. Runte for assistance with genotyping; A. Kwitek-Black for comparative mapping; and T. Kurth, K. Bork, P. Regozzi, and C. Thomas for excellent technical assistance. This work was supported by National Heart, Lung, and Blood Institute grant 1P50-HL-54998.

1 May 2001; accepted 17 October 2001

Estimation of Epidemic Size and Incubation Time Based on Age Characteristics of vCJD in the United Kingdom

Alain-Jacques Valleron,¹ Pierre-Yves Boelle,¹ Robert Will,² Jean-Yves Cesbron³

The size of the variant Creutzfeldt-Jakob Disease (vCJD) epidemic in the United Kingdom is a major public health concern and a subject of speculation. The cases are young (mean age = 28). Assuming that the risk of developing the disease in susceptible exposed subjects decreases exponentially with age after age 15, that all infections occurred between 1980 and 1989, and that the distribution of the incubation period is lognormal, we estimate that the mean duration of the incubation period is 16.7 years [95% confidence interval (CI): 12.4 to 23.2] and that the total number of cases will be 205 (upper limit of the 95% CI: 403).

As of 1 May 2001, there were 97 cases of definite ($n = 86$) or probable ($n = 11$) variant Creutzfeldt-Jakob Disease (vCJD) in the United Kingdom. These patients probably contracted the disease by oral ingestion of food contaminated by the agent of bovine spongiform encephalopathy (BSE), before the UK bovine-specified risk materials (SRM) ban in 1989 (1). The number of BSE-infected animals is estimated to have been in the range of 900,000 to 1,130,000, with between 460,000 and 482,000 slaughtered for consumption before the introduction of the November 1989 specified offal ban (2). The epidemic may have started as early as 1980 (3), and the number of people exposed to potentially infective doses through food may

be extremely high (4). Therefore, one could pessimistically assume that virtually everybody in the population has been in contact with food, or bovine products, originating from BSE-infected animals. The public health response in Europe has been to develop procedures and diagnostic tests that avoid, as far as possible, the entry of any infected animal into the food chain. The consequences of the BSE epidemic in terms of human disease are not yet known: With different assumptions for risk analysis, in 1997, the cumulative cases of vCJD in the United Kingdom were estimated from as few as 75 to as many as 80,000 (5) and more recently from 70 to 136,000 cases (6). These estimates are markedly dependent on assumptions made about the mean duration of the incubation period. Unfortunately, no studies in animals or of other human spongiform encephalopathies provide precise data for the incubation period. In addition, the observation that only a few cattle, and often only one, from the same age cohort in a herd have developed BSE suggests that additional individual or environmental factors may influence the development of the disease. It is be-

¹Epidemiology and Information Sciences, INSERM U444, CHU Saint-Antoine, Université Pierre et Marie Curie et Assistance Publique-Hôpitaux de Paris, 27 rue Chaligny, 75012 Paris, France. ²National Creutzfeldt-Jakob Disease Surveillance Unit, Western General Hospital, Edinburgh EH4 2XU, UK. ³Immunité Anti-Infectieuse JE 2236, UFR de Médecine de Grenoble, Université Joseph Fourier, Domaine de la Merci, 38706 La Tronche, France.

REPORTS

cause the relation between the risk of the disease and the dose, route of infection, host factors, and environmental factors is unknown that predictions of the possible future of the epidemic rely on mathematical extrapolations based on currently available data.

One striking epidemiological characteristic of vCJD is the young age distribution of the cases (Fig. 1A). The mean age at death of these patients is 28 years. Only 6 of 90 patients who have died were older than 50 years, in comparison with 93% of cases of sporadic CJD (7). This observation provides a clue to the incubation period in vCJD, i.e., the “age of the patient at the diagnosis” = “age at infection” + “incubation time.” The ages of the patients with vCJD at diagnosis and death are collated by the National Creutzfeldt-Jakob Disease Surveillance Unit in the United Kingdom. The causal link between BSE infection in cattle and human vCJD indicates that the age at infection in people must parallel the course of the BSE epidemic. One can therefore compute the distribution of incubation periods by a deconvolution technique taking into account the censorship of current data, as was done successfully with acquired immune deficiency syndrome (AIDS) (8). In vCJD, the analysis of the censored data is more complicated because the possible dates of infection range over an extended period of time. Thus, additional statistical modeling was necessary. The model we propose provides estimates of the incubation period and the future burden of the vCJD epidemic.

To explain that most of the cases are young, two (nonexclusive) hypotheses have to be made: (i) the incubation period is shorter in the young than in the old, and (ii) the young are more susceptible to infection. We do not prejudge the underlying determinants to these hypotheses that may be related to differences in susceptibility or other unknown factors. If the first assumption was true, older cases with longer incubation periods have a greater chance of being identified later than younger cases. Because there is no statistically significant relation between the date of diagnosis and age (Fig. 1B), we discarded this first assumption. We therefore chose the second assumption, namely that the probability per year that a susceptible subject becomes infected [the “force of infection” (9)] decreases with age a . For the calculation, we assumed that all susceptible children aged between 0.5 and 15 years, exposed to the infectious agent on year t , experience the same force of infection. After 15 years of age, we assume that the force of infection decreases exponentially (10). This hypothesis is in line with the epidemiological data, which show that vCJD rarely occurs in older subjects. The age cutoff of 15 years was an arbitrary choice. However, this parameter has been further studied in a sensitivity analysis. We assume that the whole population may have been exposed to the infectious agent, with the exception of infants

aged less than 0.5 years.

We denote by $S(t,a)$ the survival function from all causes of mortality at date t and age a , by $\lambda(t,a)$ the force of infection at date t and age a , and by h the probability density function of the duration of the incubation period. $\lambda(t,a)$ is the product of a function $f(a)$ of age and of a function $g(t)$ of calendar time: $\lambda(t,a) = f(a)g(t)$. The choices of functions f and g express the age-risk relation described above and the variation of the BSE infectious risk, respectively (11). The probability distribution of the duration of the incubation period, $h(u)$, was assumed to be lognormal, because this is a widely accepted model in infectious diseases (12). A Gamma distribution was also used in the sensitivity analysis.

The full statistical formulation necessary to write the likelihood equation implies that a probabilistic model be chosen to estimate the population at risk of developing the disease, i.e., the exposed/susceptible subjects. According to the formulation we chose, the dates and ages of the observed cases are a realization of a planar Poisson process (13). In this model, s is the average number of individuals per generation who can be infected, provided they have been exposed to a sufficient dose of infectious agent and have the appropriate individual characteristics (genetic, environmental, and so forth). In

each generation, these s subjects pass through the 1980–89 “window” at different ages, and therefore the final number of persons who will ultimately develop the disease varies accordingly. Finally, the probability $P(t,a)dadt$ that a case aged between a and $a + da$ occurs between times t and $t + dt$ is expressed mathematically, then the maximum likelihood approach is used to estimate the unknown parameters from the available observations, and the future of the epidemic is assessed with these parameter values (14).

The distribution of the vCJD incubation period that best fits the data within the framework of our model has a mean of 16.7 years, with a standard deviation of 2.6 years. The 95% upper percentile of this distribution is 21.4 years. The 95% confidence interval (CI) of the estimates of the mean and standard deviation is relatively narrow: The 95% CI for the estimate of the mean incubation period is 12.4 to 23.2 years, and the 95% CI of the standard deviation is 0.9 to 8 years (10). The decrease in susceptibility to infection in exposed subjects older than 15 years, as estimated from the parameter α , was found to be very sharp: 16% per year of age (CI: 12 to 23%). This means that, under the best fitting hypothesis, an individual aged 20 years in 1981 had 55% less risk of becoming infected than a child aged 15

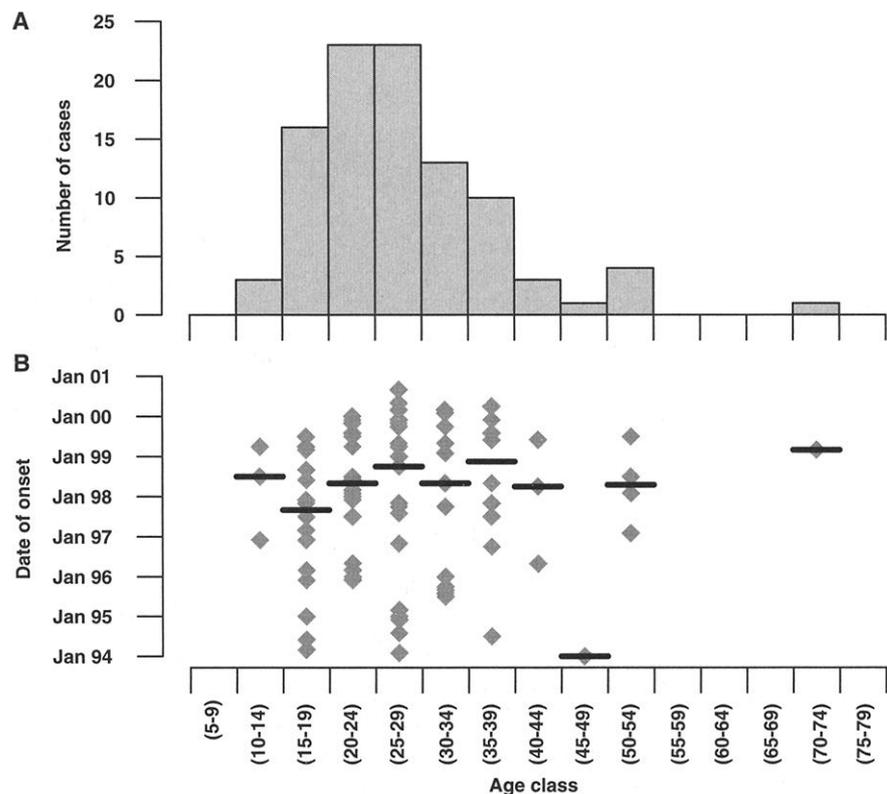


Fig. 1. (A) Age distribution of vCJD cases. Data are as of 1 May 2001, including 97 patients. (B) The distribution of the date of onset of vCJD by age class. Diamonds indicate observed values; a line is drawn at the median date. There is no apparent relation between age and date of diagnosis (correlation coefficient = 0.03, $P = 0.8$), as would have been expected if the incubation time was strongly age dependent, causing older cases to be detected later than younger cases.

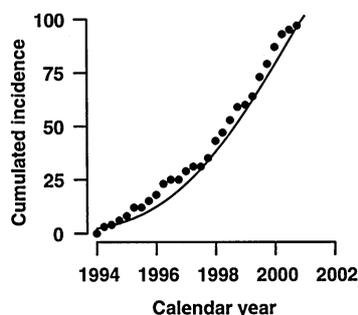


Fig. 2. Observed cumulated incidence of vCJD (dots) and incidence predicted by the model (line).

years (99.9% for an individual aged 70).

The model predictions fit the past cumulated incidences, as shown in Fig. 2. Similarly, there is a good fit to the observed age distribution (10), as the model predicts that 22 out of the first 97 diagnosed cases should be less than 20 years of age (19 in the surveillance data set) and 82 should be less than 40 years of age (87 in the surveillance data set).

Details of the variation in the risk function (14) with age and time up to 2012 are shown in Fig. 3: The model predicts that the peak of the epidemic will be in 2000/2001 and that the annual number of cases should gradually decrease after this date. The total number of expected vCJD cases is, according to our model, low at 205 cases (upper limit of the 95% CI: 403). One prediction from the model is that in the next few years, the age distribution will become bimodal and that the proportion of older patients will increase.

The model presented here addresses human infections occurring between 1980 and 1989, as we assume that the likelihood of new human infection with BSE after the bovine SRM ban in 1989 should have been considerably reduced.

The strength of the method used in this report is that it is focused on the most striking epidemiological characteristic of vCJD, namely, the age distribution of the cases. In contrast to other models, which are based primarily on scenarios related to dietary exposures to BSE in the United Kingdom, our method does not make any assumptions about this parameter that cannot currently be estimated.

To express the necessary dependence between age and force of infection, we made one of the most parsimonious choices possible, i.e., a function beginning with a plateau during childhood, followed by an exponential decrease. The sensitivity analysis in which all possible alternatives to the 15-year age limit were tested ruled out an even simpler model with only one exponential function (age limit = 0) and found as the optimal value an age limit of 16. We have no clear physiological explanation for this age limit: Although a relation with puberty may be hypothesized, experimental evidence should be sought to support such a

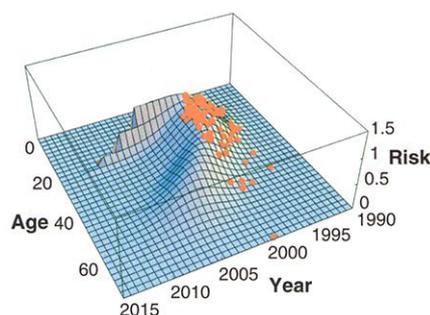


Fig. 3. Risk function for onset of vCJD as a function of age and time. Red dots indicate the 97 observed cases as of 1 May 2001.

possibility.

The results presented here are based on a lognormal distribution of the incubation period, which is commonly used in the epidemiology of infectious diseases. However, we repeated the estimation process assuming that the incubation period was Gamma distributed and found almost identical values for the incubation parameters, the future epidemic size of the epidemic, and the occurrence of a bimodal age distribution in the future (10).

The incubation time we find for vCJD is longer than in human growth hormone-related CJD, which is between 9 and 10 years (15) in the sensitive homozygous genotypes. It is shorter than the one estimated in Kuru (16), which may exceed three decades, although, a priori, one would have expected a longer value because in Kuru, there is no species barrier and the disease was transmitted orally as in vCJD. The PRNP 129 genotype has a crucial importance in determining the risk of developing CJD. Yet to date, all tested vCJD patients have been homozygous for methionine at codon 129, as are about 40% of the total UK population. If there are two subpopulations of vCJD patients, one with “short” incubation times and methionine homozygosity, the other with “long” incubation times and valine homozygosity or heterozygosity, our method would not identify the second distribution parameters, because cases of vCJD with genotypes other than methionine homozygosity have not yet been identified.

Our age-risk model allows an estimate of the future size of the epidemic. If we independently introduce our estimates of the incubation period in the Ghani or Cousens models, we find similar “low” predictions: 80 to 630 cases with the Ghani model and 801 with the Cousens model (17). In conclusion, our prediction of the epidemic of vCJD lies in the “optimistic” end of the ranges of previously published figures, and this low value is in favor of a large species barrier between cattle and humans.

References and Notes

1. M. E. Bruce *et al.*, *Nature* **389**, 498 (1997).
2. N. M. Ferguson, C. A. Donnelly, M. E. J. Woolhouse, R. M. Anderson, *Philos. Trans. R. Soc. London Ser. B Biol. Sci.* **352**, 803 (1997).

3. C. H. Cohen, A. J. Valleron, *Int. J. Epidemiol.* **28**, 526 (1999).
4. “Opinion of the scientific steering committee on the human exposure risk (HER) via food with respect to BSE” (European Union Scientific Steering Committee, 10 December 1999)
5. S. N. Cousens, E. Vynnycky, M. Zeidler, R. G. Will, P. G. Smith, *Nature* **385**, 197 (1997).
6. A. C. Ghani, N. M. Ferguson, C. A. Donnelly, R. M. Anderson, *Nature* **406**, 583 (2000).
7. Surveillance data collected by the UK Creutzfeldt-Jakob Disease Surveillance Unit are available at www.cjd.ed.ac.uk/.
8. S. Chevret, D. Costagliola, J. J. Lefrere, A. J. Valleron, *J. Epidemiol. Community Health* **46**, 582 (1992).
9. The “force of infection” is defined in mathematical epidemiology to quantify the risk per unit of time (usually, the year) that a susceptible individual has to acquire the disease after being exposed to the infectious agent. For a complete review of the concept of force of infection, see R. M. Anderson and R. May [*Infectious Diseases of Humans: Dynamics and Control* (Oxford Science Publications, Oxford, 1991), chaps. 3 and 8].
10. Supplemental materials are available at *Science Online* at www.sciencemag.org/cgi/content/full/294/5547/1726/DC1.
11. The function $f(a)$ was taken in the form $f(a) = 0$ if $a < 0.5$, as infants younger than 0.5 year were assumed to not be exposed to the BSE agent; $f(a) = \lambda_0$ if $0.5 < a < 15$, which corresponds to the constant risk assumed in children before 15; and $f(a) = \lambda_0 e^{-\alpha(a-0.5)}$ if $a > 15$. The function $g(t)$ was chosen to reflect the increasing probability of the presence of contaminated meat products in the human diet, in parallel to the increasing incidence of BSE in animals over the period 1980–89. Possible exposures after 1989 were not considered in this analysis. Mimicking the incidence profile observed in the BSE epidemic during the same period, we set $g(t) = \gamma_0 e^{-\gamma(t-1989)}$. From the knowledge of the doubling time observed in the BSE epidemic (14 months) during the 1980–88 period (2), we extrapolate $\gamma = \ln(2)/(14/12) = 0.596$. As only the product $\lambda_0 \gamma_0$ was identifiable, because of the form for the force of infection, we fixed γ_0 at 1.
12. P. E. Sartwell, *Am. J. Hyg.* **51**, 310 (1950).
13. D. R. Brillinger, *Biometrics* **42**, 693 (1986).
14.
$$P(t,a)dt da = S(t,a) \int_0^a \left\{ \exp \left[- \int_0^v \lambda(t-u, u) du \right] \lambda(t-v, v) h(a-v) dv \right\} dt da$$

The log-likelihood of the observed cases is

$$\sum_i \log [sP(t_i, a_i)] - \int \int_D sP(t, a) dadt,$$

where the first sum is over all observed cases and the integration is made in the domain D corresponding to ages between 0.5 and 100 years and to years between 1980 and 2000. s is the intensity of the Poisson birth process. Numerical maximization of the likelihood was done with a modified Levenberg Marquardt algorithm (SLATEC) after reparameterization of the model. A logarithmic transformation was used for all variables but incubation mean time. The maximization was independently performed with a Newton-Raphson method (NAG), and all presented results were in agreement in the two methods, with differences between estimates of less than 10^{-2} . Numerical integrations were carried out to a precision of 10^{-4} .

15. J. Huillard d’Aignaux *et al.*, *Neurology* **53**, 1197 (1999).
16. R. L. Klitzman, M. P. Alpers, D. C. Gajdusek, *Neuroepidemiology* **3**, 20 (1984).
17. The figure obtained from the Cousens model when an incubation period of mean 15 years and 95th percentile of 22 years are assumed.
18. We thank J. Gagnon for helpful discussions.

4 October 2001; accepted 20 October 2001